# Getting fit for the Midterm! 

Econ 140, Section 4

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## Roadmap

1. Recap
2. Interaction terms (Q6)
3. Logs (Q4)
4. Topics we've glossed over so far

## Any questions?

... Some comments on the evaluations asked for more space to answer left-over questions from the lecture: Now is the time!

Recap

## Recap: OVB (Very important!)

We can summarize everything of OVB in three equations. Let $Y_{i}$ be the outcome variable, $X_{i}$ our regressor of interest, and $Z_{i}$ the "omitted" variable.

$$
\begin{aligned}
\text { [Long regression] } & Y_{i}=c_{1}+\beta_{L} X_{i}+\delta Z_{i}+e_{i} \\
\text { [Short regression] } & Y_{i}=c_{2}+\beta_{S} X_{i}+u_{i} \\
\text { [Auxiliary regression] } & Z_{i}=c_{3}+\gamma X_{i}+v_{i}
\end{aligned}
$$

Then, the Omitted variable bias formula states that:

$$
\underbrace{\beta_{S}}_{\text {Short }}=\underbrace{\beta_{L}}_{\text {Long }+}+\underbrace{\delta}_{\text {Omitted } \times} \cdot \underbrace{\gamma}_{\text {Included }}
$$

We call $\delta \gamma$ the omitted variable bias. We can appraise the direction of the bias by multiplying our guesses for the signs of $\delta$ and $\gamma$. If either $\delta=0$ or $\gamma=0$, then OVB is zero!

## Recap: Understanding bias in OLS regressions

- We can use the OLS formula to understand how bias works in OLS regression

$$
\hat{\beta}_{1}=\frac{\operatorname{Cov}\left(X_{i}, Y_{i}\right)}{\operatorname{Var}\left(X_{i}\right)}
$$

- For OVB: We know the true $Y_{i}$ and plug it in
- For measurement error: We know what $X_{i}$ and plug it in
- Simplify using the following rules:

$$
\begin{aligned}
& \operatorname{Var}(X+Y)=\operatorname{Var}(X)+\operatorname{Var}(Y)+2 \operatorname{Cov}(X, Y) \\
& \operatorname{Var}(X-Y)=\operatorname{Var}(X)+\operatorname{Var}(Y)-2 \operatorname{Cov}(X, Y) \\
& \operatorname{Cov}(a X+b Y, Z)=a \operatorname{Cov}(X, Z)+b \operatorname{Cov}(Y, Z) \\
& \operatorname{Cov}(X, X)=\operatorname{Var}(X) \\
& \operatorname{Var}(a X)=a^{2} \operatorname{Var}(X) \\
& \operatorname{Cov}(X, Y)=0, \text { if } X \text { and } Y \text { are independent. } \\
& \operatorname{Var}(X) \geq 0 .
\end{aligned}
$$

## Recap: Making OLS more interesting

- We saw that we can extend the simple OLS framework

$$
Y_{i}=\beta_{0}+\beta_{1} X_{i}+e_{i}
$$

to something richer:

$$
Y_{i}=\beta_{0}+\beta_{1} X_{i}+\beta_{2} X_{i}^{2}+e_{i}
$$

- We will get to know many more versions of this today
- All questions of the type "how is $Y_{i}$ expected to change if we change $X_{i}$ " can be solved with partial derivatives - in this case:

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\frac{\partial Y_{i}}{\partial X_{i}}=
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- All questions of the type "how is $Y_{i}$ expected to change if we change $X_{i}$ " can be solved with partial derivatives - in this case:

$$
\frac{\partial Y_{i}}{\partial X_{i}}=\beta_{1}+2 \cdot \beta_{2} \cdot X_{i}
$$

Interaction terms (Q6)

## Interaction terms: Making OLS more interesting (Q6)

Let us consider the model

$$
Y_{i}=\beta_{0}+\beta_{1} X_{1 i}+\beta_{2} X_{2 i}+e_{i}
$$

where $Y_{i}$ is a country's GDP per capita, $X_{1 i}$ the value of its natural resources, and $X_{2 i}$ a measure of how democratic it is.

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Keeping democracy fixed, increasing the value of a country's natural resources by one unit is associated with $\beta_{1}$ higher GDP per capita.
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2. How do we interpret $\beta_{2}$ ?

Keeping natural resources fixed, increasing a country's democracy score by one unit is associated with $\beta_{2}$ higher GDP per capita.

## Interaction Terms (ii) (Q6)

Now, let us extend the model to:

$$
Y_{i}=\beta_{0}+\beta_{1} X_{1 i}+\beta_{2} X_{2 i}+\beta_{3} X_{1 i} X_{2 i}+e_{i}
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1. What is the "effect" of $X_{1 i}$ on $Y_{i}$ ?

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4. How do we interpret $\beta_{1}+\beta_{3}$ ?

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3. How do we interpret $\beta_{2}$ ? The effect of an additional unit of $X_{2 i}$, if $X_{1 i}$ is equal to 0 .
4. How do we interpret $\beta_{1}+\beta_{3}$ ? The effect of an additional unit of $X_{1 i}$, if $X_{2 i}$ is equal to 1 .

Rule of thumb: Always use partial derivatives to make sure that you are right!

## Practice exam question: 1a)

The Ministry of Truth is interested in a rumour that air pollution could impact mental health. One of the most harmful pollutants is fine particulate matter PM2.5, which comes from operations that involve the burning of fuels such as wood, oil, coal, etc. A research team is sent to investigate the rumour. The team randomly selects and surveys 19,920 people across 71 districts of the country. The key variable, Exposure $E_{i}$, is a dummy variable equal to 1 if the individual $i$ is exposed to a large amount of PM2.5 in the last 24 hours, and 0 otherwise. The team also conducts a standardised questionnaire to record depressive symptoms in the last month, called the Kessler Psychological Distress scale (K6). The questionnaire results in a score, Depression ${ }_{i}$, that ranges from 0 to 24; and the higher the score, the more severe the depressive symptoms for individual $i$. The variable has a sample average of 2.96 . Running regressions with Depression $D_{i}$ as the dependent variable, you obtain the following results:

## Practice exam question: 1a)

Dependent variable: Depression i

| Regressor | $(1)$ | $(2)$ | $(3)$ |
| :--- | :---: | :---: | :---: |
| Exposure ${ }_{i}$ | 0.834 | 0.614 | 0.554 |
|  | $(0.032)$ | $(0.045)$ | $(0.042)$ |
| Exposure $_{i} \times$ Female $_{i}$ |  | 0.065 |  |
|  |  | $(0.024)$ |  |
| Female $_{i}$ |  | -0.739 | -0.825 |
|  |  | $(0.036)$ | $(0.066)$ |
| Age $_{i}$ |  |  | 0.452 |
|  |  |  | $(0.132)$ |
| Age $_{i}^{2}$ |  |  | 0.524 |
|  |  |  | $(0.121)$ |

Notes: All estimations contain a constant term. Robust standard errors are in the parentheses. Age ${ }_{i}$ is the age (years old) of individual $i$, and Age $_{i}^{2}$ is the square of Age $_{i}$.

## Pratice Exam question: 1a)

a) Interpreting the coefficient in Column (1), a journalist, Katherine, claims: "Since participants are randomly selected, we can infer that exposure to a large amount of PM2.5 does cause depression."
i. Explain carefully why Katherine is wrong, specifying the direction of bias(es) if there is any. Which assumption(s) would she need to impose for the causality claim to hold?
ii. What is the correct interpretation from Column (1) that Katherine should have made?

## (Detailed) Suggested Answer: 1a)

i. Random selection is not the same thing as random assignment to treatment! Survey respondents may be systematically different from each other in ways that are correlated with depression and pollution exposure. Therefore, the results from a regression can not be interpreted causally (and are biased). A priori, it is unclear in which direction the bias would go, but we could imagine that (Only one explanation needed for exam):
On rainy days, pollution is lower (-) and people may be reporting more depression symptoms (+), leading to downward bias. More wealthy people choose to live in less polluted areas (-) and they may have less depression (e.g., better access to mental health resources) (-), leading to upward bias. For the regression causality claim to hold, we need to assume that: People exposed to pollution and those not exposed to pollution would have, on average, the same depression level, had they been exposed to the same level of pollution. In other words: Both groups would have to have the same potential depression outcomes.
ii. On average, people that were exposed to pollution had a 0.8 points higher score on the depression scale. The difference between the two groups is significant at the 5\% level.

## Pratice Exam question: 1b)

b) Interpret column (2) of the regression table
i. A colleague notes the the coefficient on Female ${ }_{i}$ is significant, and states: "The effect of being female on depression is significantly different from zero". Do you agree with the statement? Why or why not?
ii. How is pollution exposure related to depression, for men?

And how for women?

## (Detailed) Suggested Answer: 1b)

i. It is difficult to make such interpretations when interaction terms are involved. Taking partial derivatives, the "effect" of being female is:

$$
\frac{\partial \text { Depression }_{i}}{\partial \text { Female }_{i}}=-0.739+0.065 \cdot \text { Exposure }_{i}
$$

We can do inference (and test significance) at Exposure ${ }_{i}=0$ (just looking at the coefficient for female, -0.739 is significant). We can also do it for any other level of exposure, but for that we also need to take the other coefficient into account and cannot just use the table.
Comment 1: Less relevant for the exam, but important to be aware of!
Comment 2: We can always do inference on the interaction term, which is significant here!
ii. Taking partial derivatives, the "effect" of pollution exposure is:

$$
\frac{\text { Depression }_{i}}{\partial \text { Exposure }_{i}}=0.614+0.065 \cdot \text { Female }_{i}
$$

Hence, the "effect" for Males is 0.614 and the "effect" for Females is larger $(0.614+0.065=0.779)$. The effect of pollution is also signficantly larger than for men, because the coefficient on the interaction term is significantly different from zero.
Logs (Q4)

## Notes on logarithms (Q4)

- We can take logs of whole equations to get linear models (problem set)
- We can also take logs of specific variables, especially when they have long tails (wealth in the US, GDP per capita, etc.)
- We can get to the right interpretation of log-specifications with just math


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- We can take logs of whole equations to get linear models (problem set)
- We can also take logs of specific variables, especially when they have long tails (wealth in the US, GDP per capita, etc.)
- We can get to the right interpretation of log-specifications with just math
- But I will make your life easier with a cheat sheet.


## Logs: Cheatsheet (Wooldridge version) (Q4)

Summary of Functional Forms Involving Logarithms

| Model | LHS | RHS | Interpretation of $\boldsymbol{\beta}_{1}$ |
| :--- | :---: | :---: | :---: |
| Level-level | $y$ | $x$ | $\Delta y=\beta_{1} \Delta x$ |
| Level-log | $y$ | $\log (x)$ | $\Delta y=\left(\beta_{1} / 100\right) \% \Delta x$ |
| Log-level | $\log (y)$ | $x$ | $\% \Delta y=\left(100 \beta_{1}\right) \Delta x$ |
| Log-log | $\log (y)$ | $\log (x)$ | $\% \Delta y=\beta_{1} \% \Delta x$ |

Table taken from Wooldridge (2011)

## Logs: Cheatsheet II (Jonathan's version) (Q4)

| Model | LHS | RHS | A change in <br> $x$ by $\ldots$ | is associated <br> with a change <br> in $y$ by $\ldots$ |
| :--- | :---: | :---: | :---: | :--- |
| Level-Level | $y$ | $x$ | 1 unit | $\beta_{1}$ units |
| Level-Log | $y$ | $\log (x)$ | $1 \%$ | $\beta_{1} / 100$ units |
| Log-Level | $\log (y)$ | $x$ | 1 unit | $100 \beta_{1} \%$ |
| Log-Log | $\log (y)$ | $\log (x)$ | $1 \%$ | $\beta_{1} \%$ |

If you want to get a bonus star from me, write "approximately" in log-interpretations.

Topics we've glossed over so far

## Any questions?

... Some comments on the evaluations asked for more space to answer left-over questions from the lecture: Now is the time!

## Bad controls

- Not all controls are good controls
- Some controls are called "bad controls". These are:

1. Variables that are themselves outcomes of a treatment: What happens if you control for the change in English test scores in the regression below?

|  | Treatment | Control |
| :--- | :--- | :--- |
| Change in Math Scores | 2 | 1 |
| Change in English Scores | 2 | 1 |

2. Variables that moderate the treatment effect, e.g. controlling for occupation choice in gender wage gap regression...

- Rule of Thumb: Good controls are either pre-determined or immutable characteristics.
- Another way to think about it: Controls help us make "apples to apples" comparisons. Which apples matter?


## What if the outcome variable is binary (a dummy variable)?

Let's run the regression

$$
\text { Defaulted }_{i}=\alpha+\beta \text { Credit̃ Score }_{i}+e_{i}
$$

where Defaulted $_{i}$ is equal to 1 if individual $i$ has ever defaulted on a loan (mortgage, credit card, auto loan, etc.), and Credit̃ Score is i's credit score, minus the average credit score in the sample (Note: US credit scores range from 300 to 850 points).

1. You run a regression and get $\hat{\alpha}=0.1$. How do you interpret this? Does this number make sense here?
2. Your estimate for $\beta$ is $\hat{\beta}=0.001$. Interpret.

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With a dummy dependent variable, changing $X_{i}$ by one unit increases the probability of $Y_{i}=1$ by $\hat{\beta} \cdot 100$ percentage points.

## Inference and the variance of $\hat{\beta}_{0 L S}$ (Q3)

The variance of the OLS estimator is $\operatorname{Var}\left(\hat{\beta}_{1}^{0 L S}\right)=\frac{\sigma_{\epsilon}^{2}}{N \cdot \operatorname{Var}\left(X_{i}\right)}$. We expect to get more precise estimates if

- The variance of $X_{i}$ increases
- The variance of the error term $\epsilon_{i}$ decreases
- The sample size $N$ increases


## Hypothesis testing

$$
\begin{aligned}
& \left|\frac{\hat{\beta}}{\operatorname{SE}(\hat{\beta})}\right| \geq 1.96 \\
& \Leftrightarrow \mid \mathrm{t} \text {-stat } \mid \geq 1.96 \\
& \Leftrightarrow \text { p-value } \leq 0.05
\end{aligned}
$$

If you are testing the null hypothesis $H_{0}: \beta=0$, then all of these are equivalent, and you can use any of these.

